

Finite Dimensional Statistical Inference

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Introduction

Deconvolution

Given \mathbf{A} , \mathbf{B} $n \times n$ independent hermitian random matrices:

① **Additive free deconvolution:**

To derive the eigenvalue distribution of \mathbf{A} (denoted $\lambda(\mathbf{A})$) knowing $\lambda(\mathbf{A} + \mathbf{B})$ and $\lambda(\mathbf{B})$.

② **Multiplicative free deconvolution:**

To derive the eigenvalue distribution $\lambda(\mathbf{A})$ knowing $\lambda(\mathbf{AB})$ and $\lambda(\mathbf{B})$.

Previous Works

- $n = 1$ (Scalar Case - Classical Probability)
- $n \rightarrow \infty$ (Asymptotic Case - Free Probability)

Previous Works

- $n = 1$ (Scalar Case - Classical Probability)
- $n \rightarrow \infty$ (Asymptotic Case - Free Probability)

What happens between the Asymptotic Case and the Scalar Case?

Asymptotic Case

-Analytical Method

Stieltjes transform of a probability measure μ

$$s_{\mu}(z) = \int \frac{1}{\lambda - z} dF^{\mu}(\lambda)$$

-Moment Method

p -th moment of \mathbf{A}

$$t_{\mathbf{A}}^{n,p} = \mathbb{E} [\text{tr}(\mathbf{A}^p)] = \int \lambda^p d\rho(\lambda)$$

Analytical Method

- ✓ Based on the Stieltjes transform
- ✓ Works for any measures
- ✗ Can not be implemented

Moment Method

- ✓ Based on the Moments
- ✗ Works only for measures with ALL its moments
- ✓ Can be implemented

Objective of the presentation

To propose algorithmic methods to compute
deconvolution for finite size matrices.

In particular, we focus on Gaussian matrices.

Model 1: Multiplicative model

$$\mathbf{Y} = \mathbf{D}\mathbf{X}$$

where

- where \mathbf{X} is an $n \times N$ complex standard Gaussian matrix (i.e. with i.i.d., mean zero, variance 1, complex, Gaussian entries),
- \mathbf{D} is an $n \times n$ deterministic matrix.

Multiplicative Deconvolution intends to express the moments of $\mathbf{D}\mathbf{D}^H$ based only on the moments of $\mathbf{Y}\mathbf{Y}^H$ (known as [One side correlated zero mean Wishart matrix](#)).

Model 2: Additive Model

$$\mathbf{Y} = \mathbf{D} + \mathbf{X}$$

where

- \mathbf{X} is an $n \times N$ complex standard Gaussian matrix;
- \mathbf{D} is an $n \times N$ deterministic matrix.

Additive Deconvolution intends to express the moments of $\mathbf{D}\mathbf{D}^H$ based only on the moments of $\mathbf{Y}\mathbf{Y}^H$ (known as [Non-central Wishart matrix](#)).

Main Results

Theorem 1:

Let n, N be positive integers, \mathbf{X} be complex, standard, Gaussian and \mathbf{D} a (deterministic) $n \times n$ matrix. For any positive integer p , we have

$$\mathbb{E} \left[\text{tr} \left(\left(\mathbf{D} \frac{1}{N} \mathbf{X} \mathbf{X}^H \right)^p \right) \right] = \sum_{\pi \in \mathcal{S}_p} N^{k(\hat{\pi}) - p} n^{l(\hat{\pi}) - 1} D_{\hat{\pi}|\text{odd}},$$

where $\hat{\pi}|\text{odd}$ is the partition consisting of the equivalence classes/blocks of $\hat{\pi}$ which are contained within the odd numbers; $D_\rho = \prod_{i=1}^k \text{tr}(\mathbf{D}^{|\rho_i|})$ whenever $\rho = \{\rho_1, \dots, \rho_k\}$ is a partition with blocks ρ_i ; and $|\rho_i|$ is the number of elements in ρ_i .

Denoted $M_p = \mathbb{E} \left[\text{tr} \left(\left(\mathbf{D} \frac{1}{N} \mathbf{X} \mathbf{X}^H \right)^p \right) \right]$, $D_p = \mathbb{E} [\text{tr}(\mathbf{D}^p)]$ and $c = \frac{n}{N}$

$$M_1 = D_1$$

$$M_2 = D_2 + cD_1^2$$

$$M_3 = \left(1 + \frac{1}{N^2} \right) D_3 + 3cD_2D_1 + c^2D_1^3$$

\vdots

By a simple recursion, we can express D_p from M_p . For the first three moments these recursions become

$$D_1 = M_1$$

$$D_2 = M_2 - cM_1^2$$

$$D_3 = [M_3 - 3c(M_2 - cM_1^2)M_1 + c^2M_1^3] / \left(1 + \frac{1}{N^2}\right)$$

⋮

Theorem 2:

Let \mathbf{X} be an $n \times N$ complex, standard Gaussian matrix, \mathbf{D} a deterministic $n \times N$ matrix, and set $D_p = \text{tr} \left(\left(\frac{1}{N} \mathbf{D} \mathbf{D}^H \right)^p \right)$. Also, for $q \in S_{\rho_2}$, define π by the requirement that $\pi(\rho_1) = q(\rho_2)$. We have that

$$\begin{aligned} & \mathbb{E} \left[\text{tr} \left(\left(\frac{1}{N} (\mathbf{D} + \mathbf{X})(\mathbf{D} + \mathbf{X})^H \right)^p \right) \right] \\ &= \sum_{\rho_1 \subset \{1, \dots, p\}} \frac{1}{nN^{|\rho_1|}} \sum_{\rho_2 \subset \{1, \dots, p\}} \sum_{q \in S_{\rho_2}} n^{k(\hat{\pi}) - kd(\hat{\pi})} \times \\ & \qquad \qquad \qquad N^{l(\hat{\pi}) - ld(\hat{\pi})} n^{|\sigma|} \prod_i D_{|\sigma_i|/2} \end{aligned}$$

Defining $M_p = \mathbb{E} \left[\text{tr} \left(\left(\frac{1}{N} (\mathbf{D} + \mathbf{X})(\mathbf{D} + \mathbf{X})^H \right)^p \right) \right]$, $D_p = \mathbb{E} \left[\text{tr} \left(\frac{1}{N} (\mathbf{D} \mathbf{D}^H)^p \right) \right]$ and $c = \frac{n}{N}$

$$M_1 = D_1 + 1$$

$$M_2 = D_2 + (2 + 2c) D_1 + (1 + c)$$

$$\vdots$$

- $\mathbb{E} \left[\text{tr} \left(\left(\frac{1}{\sqrt{N}} (\mathbf{D} + \mathbf{X}) \right)^p \right) \right]$, where \mathbf{X} is selfadjoint standard Gaussian

- $\mathbb{E} \left[\text{tr} \left(\left(\frac{1}{\sqrt{N}} \mathbf{D} \mathbf{X} \right)^p \right) \right]$, where \mathbf{X} is selfadjoint standard Gaussian

- $$\mathbb{E} \left[\text{tr}_n \left(\left(\mathbf{D} \frac{1}{N_1} \mathbf{X}_1 \mathbf{X}_1^H \frac{1}{N_2} \mathbf{X}_2 \mathbf{X}_2^H \cdots \frac{1}{N_k} \mathbf{X}_k \mathbf{X}_k^H \right)^p \right) \right],$$

where $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_k$ are independent complex standard Gaussian $n \times N_1, n \times N_2, \dots, n \times N_k$

- $$\mathbb{E} \left[\text{tr}_n \left(\left(\mathbf{D} \frac{1}{\sqrt{N}} \mathbf{X}_1 \frac{1}{\sqrt{N}} \mathbf{X}_2 \cdots \frac{1}{\sqrt{N}} \mathbf{X}_k \right)^p \right) \right],$$

$\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_k$ are independent selfadjoint standard Gaussian $N \times N$ matrices

Why do we study the Finite Case?

- Finite dimensional results are not strictly needed for some random matrix models
 - ⇒ Asymptotic results CAN also be applied

- Finite dimensional results are needed for some random matrix patterns
 - ⇒ Asymptotic results CAN NOT be applied
 - ⇒ We are not able to combine a set of observations into a larger compound matrix

How we can stacking the observations?

L observations of the matrix \mathbf{R}

- Horizontal

$$\mathbf{R} = [\mathbf{R}_{11}\mathbf{R}_{12}\cdots\mathbf{R}_{1L}]$$

- Vertical

$$\mathbf{R} = \begin{bmatrix} \mathbf{R}_{11} \\ \mathbf{R}_{21} \\ \vdots \\ \mathbf{R}_{L1} \end{bmatrix}.$$

- Quadratic

$$\mathbf{R} = \mathbf{D} = \begin{bmatrix} \mathbf{R}_{11} & \mathbf{R}_{1L} \\ \mathbf{R}_{2L} & \mathbf{R}_{LL} \end{bmatrix}.$$

Example

$$\mathbf{Y} = \mathbf{D}\mathbf{X}_1 + \mathbf{X}_2$$

where

- \mathbf{X}_1 is a $m \times L$ matrix
- \mathbf{X}_2 is a $n \times N$ matrix
- \mathbf{D} is $n \times m$ matrix

L independent observations $\mathbf{Y}_k = \mathbf{D}\mathbf{X}_{1,k} + \mathbf{X}_{2,k}$ ($1 \leq k \leq L$)

we can write

$$\mathbf{Y}_{1\dots L} = \mathbf{D}\mathbf{X}_{1,1\dots L} + \mathbf{X}_{2,1\dots L}$$

where

$$\mathbf{Y}_{1\dots L} = [\mathbf{Y}_1 \mathbf{Y}_2 \cdots \mathbf{Y}_L]$$

$$\mathbf{X}_{1,1\dots L} = [\mathbf{X}_{1,1} \mathbf{X}_{1,2} \cdots \mathbf{X}_{1,L}]$$

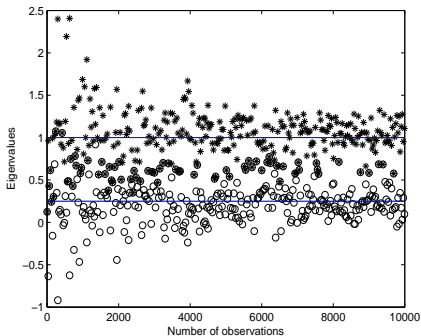
$$\mathbf{X}_{2,1\dots L} = [\mathbf{X}_{2,1} \mathbf{X}_{2,2} \cdots \mathbf{X}_{2,L}]$$

Taking

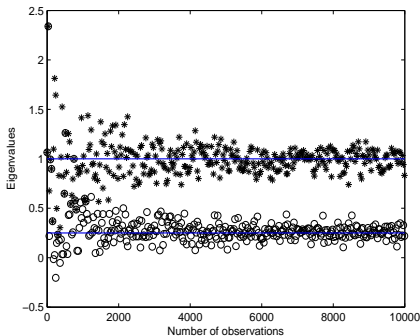
- $\mathbf{X}_1, \mathbf{X}_2$ are independent complex standard Gaussian matrices



$$\mathbf{D} = \begin{pmatrix} 1 & 0 \\ 0 & 0.5 \end{pmatrix}.$$

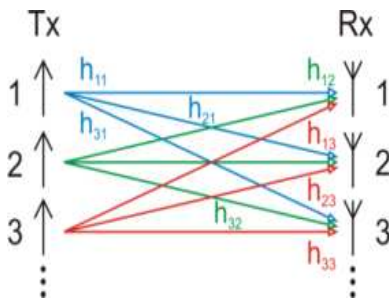


Averaging of moments



Horizontal Stacking

Rate Estimation

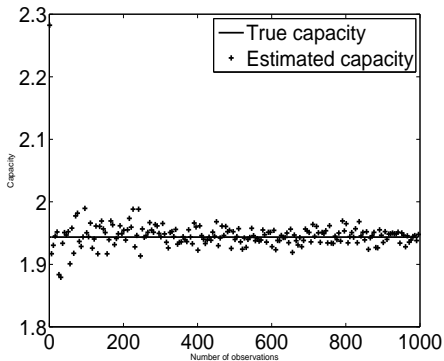
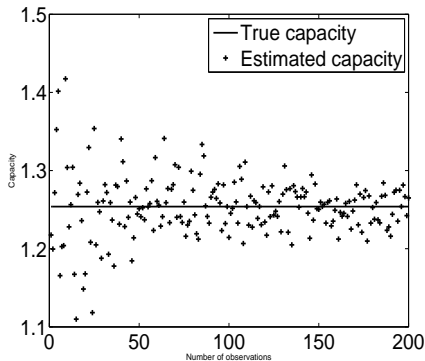


$$\mathbf{Y}_i = \mathbf{D} + \sigma \mathbf{N}_i \quad (i = 1, \dots, M)$$

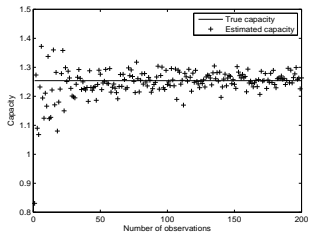
where

- \mathbf{D} is an $n \times N$ deterministic channel matrix;
- \mathbf{N}_i is an $n \times N$ standard Gaussian matrix representing the noise, with variance σ

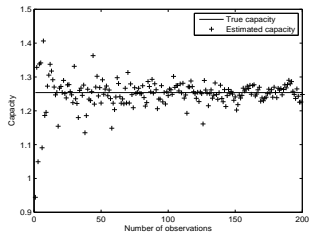
$$C = \frac{1}{n} \log_2 \det \left(\mathbf{I}_n + \frac{\rho}{N} \mathbf{D} \mathbf{D}^H \right) = \frac{1}{n} \log_2 \left(\prod_{i=1}^n (1 + \rho \lambda_i) \right)$$



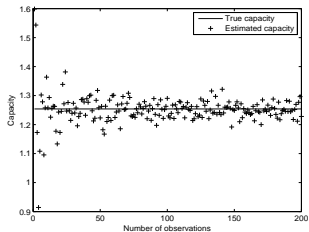
It is seen that less observations are needed for precise channel capacity estimation in the case of the 4×4 -matrix.



Quadratic Stacking

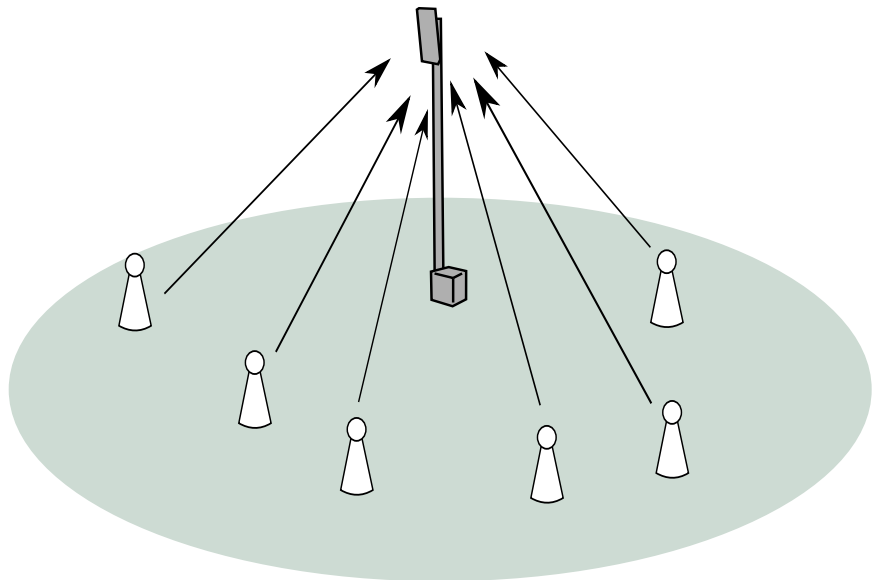


Horizontal Stacking



Averaging of Observations

Power Estimation

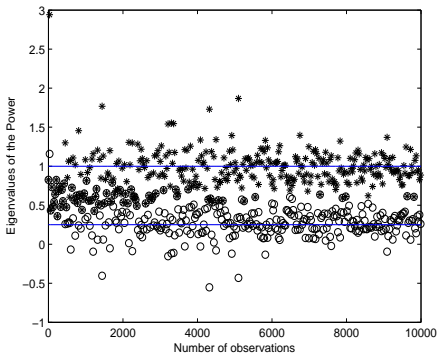


$$\mathbf{Y} = \mathbf{W}\mathbf{P}^{\frac{1}{2}}\mathbf{S} + \sigma\mathbf{N}$$

where

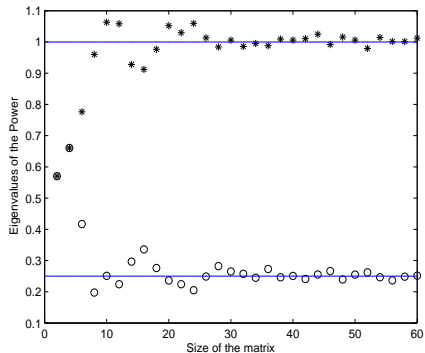
- \mathbf{W} is $N \times K$ channel gain matrix,
- \mathbf{P} is the $K \times K$ diagonal power matrix due to the different distances from which the users emit,
- \mathbf{S} is the $K \times M$ matrix of signals
- \mathbf{N} is the $N \times M$ matrix representing the noise with variance σ .

\mathbf{W} , \mathbf{S} , \mathbf{N} are independent, complex, standard Gaussian matrices.

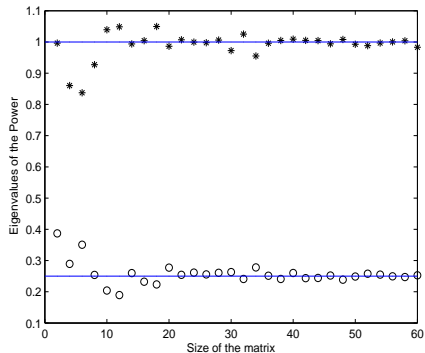


Variance of the Moments Esimator

$K = M = N$ is increased



Observations $L = 15$



Observations $L = 50$

Conclusions

- We proposed a framework to compute the moments of many types of combinations of independent Gaussian and Wishart random matrices;
- We presented a Software Implementation
- We found the optimal way of stacking the observations for stackable models;
- We solved examples of non-stackable models where asymptotic results are not applicable.